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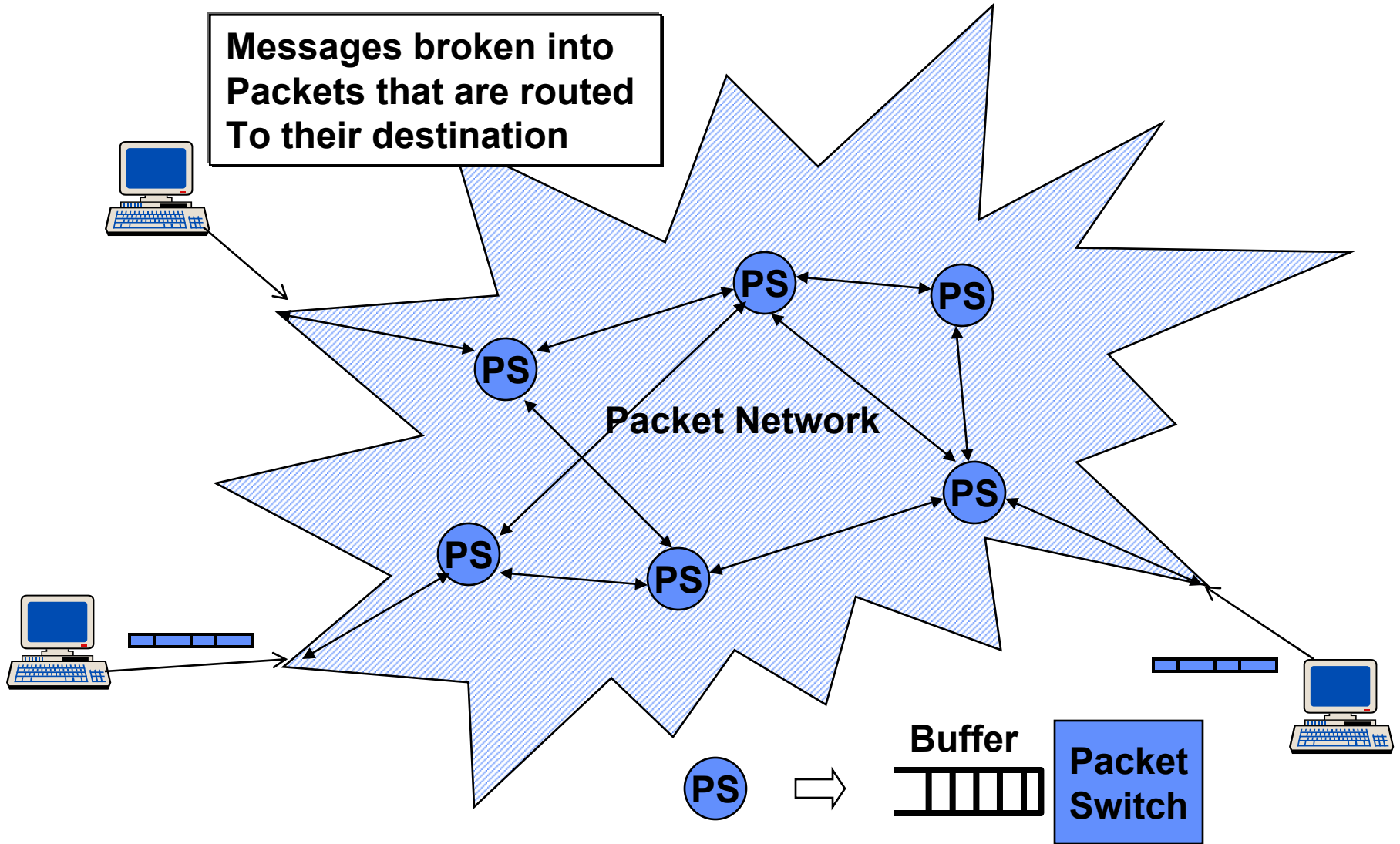
# Lectures 5 & 6

6.263/16.37

## Introduction to Queueing Theory

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# Packet Switched Networks



# Queueing Systems

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- **Used for analyzing network performance**
- **In packet networks, events are random**
  - Random packet arrivals
  - Random packet lengths
- **While at the physical layer we were concerned with bit-error-rate, at the network layer we care about delays**
  - How long does a packet spend waiting in buffers ?
  - How large are the buffers ?
- **In circuit switched networks want to know call blocking probability**
  - How many circuits do we need to limit the blocking probability?

# Random events

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- **Arrival process**
  - Packets arrive according to a random process
  - Typically the arrival process is modeled as Poisson
- **The Poisson process**
  - Arrival rate of  $\lambda$  packets per second
  - Over a small interval  $\delta$ ,

$$P(\text{exactly one arrival}) = \lambda\delta + o(\delta)$$

$$P(0 \text{ arrivals}) = 1 - \lambda\delta + o(\delta)$$

$$P(\text{more than one arrival}) = o(\delta)$$

Where  $o(\delta)/\delta \rightarrow 0$  as  $\delta \rightarrow 0$ .

- It can be shown that:

$$P(n \text{ arrivals in interval } T) = \frac{(\lambda T)^n e^{-\lambda T}}{n!}$$

# The Poisson Process

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$$P(n \text{ arrivals in interval } T) = \frac{(\lambda T)^n e^{-\lambda T}}{n!}$$

**n = number of arrivals in T**

**It can be shown that,**

$$\mathbf{E[n]} = \lambda T$$

$$\mathbf{E[n^2]} = \lambda T + (\lambda T)^2$$

$$\sigma^2 = \mathbf{E}[(n - E[n])^2] = \mathbf{E}[n^2] - E[n]^2 = \lambda T$$

# Inter-arrival times

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- Time that elapses between arrivals (IA)

$$\begin{aligned}P(\text{IA} \leq t) &= 1 - P(\text{IA} > t) \\ &= 1 - P(0 \text{ arrivals in time } t) \\ &= 1 - e^{-\lambda t}\end{aligned}$$

- This is known as the exponential distribution
  - Inter-arrival CDF =  $F_{\text{IA}}(t) = 1 - e^{-\lambda t}$
  - Inter-arrival PDF =  $d/dt F_{\text{IA}}(t) = \lambda e^{-\lambda t}$
- The exponential distribution is often used to model the service times (i.e., the packet length distribution)

# Markov property (Memoryless)

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$$P(T \leq t_0 + t | T > t_0) = P(T \leq t)$$

Proof:

$$\begin{aligned} P(T \leq t_0 + t | T > t_0) &= \frac{P(t_0 < T \leq t_0 + t)}{P(T > t_0)} \\ &= \frac{\int_{t_0}^{t_0+t} \lambda e^{-\lambda t} dt}{\int_{t_0}^{\infty} \lambda e^{-\lambda t} dt} = \frac{-e^{-\lambda t} \Big|_{t_0}^{t_0+t}}{-e^{-\lambda t} \Big|_{t_0}^{\infty}} = \frac{-e^{-\lambda(t+t_0)} + e^{-\lambda(t_0)}}{e^{-\lambda(t_0)}} \\ &= 1 - e^{-\lambda t} = P(T \leq t) \end{aligned}$$

- **Previous history does not help in predicting the future!**
- **Distribution of the time until the next arrival is independent of when the last arrival occurred!**

# Example

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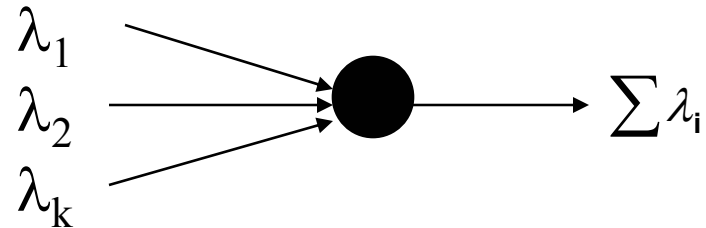
- **Suppose a train arrives at a station according to a Poisson process with average inter-arrival time of 20 minutes**
- **When a customer arrives at the station the average amount of time until the next arrival is 20 minutes**
  - **Regardless of when the previous train arrived**
- **The average amount of time since the last departure is 20 minutes!**
- **Paradox: If an average of 20 minutes passed since the last train arrived and an average of 20 minutes until the next train, then an average of 40 minutes will elapse between trains**
  - **But we assumed an average inter-arrival time of 20 minutes!**
  - **What happened?**



# Properties of the Poisson process

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- **Merging Property**

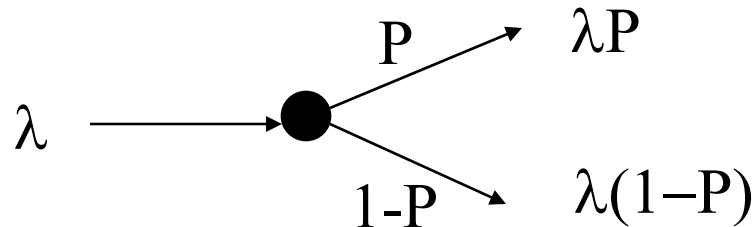


Let  $A_1, A_2, \dots, A_k$  be independent Poisson Processes  
of rate  $\lambda_1, \lambda_2, \dots, \lambda_k$

$$A = \sum A_i \text{ is also Poisson of rate } = \sum \lambda_i$$

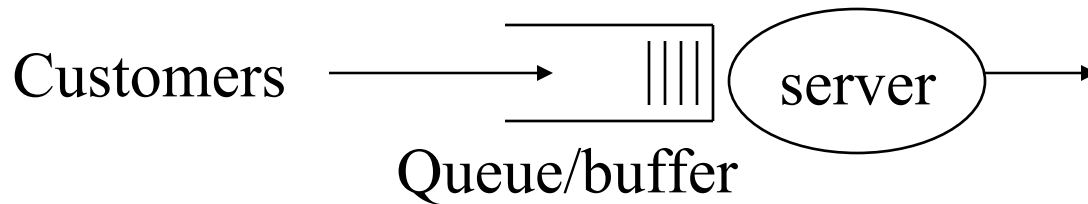
- **Splitting property**

- Suppose that every arrival is randomly routed with probability  $P$  to stream 1 and  $(1-P)$  to stream 2
- Streams 1 and 2 are Poisson of rates  $P\lambda$  and  $(1-P)\lambda$  respectively



# Queueing Models

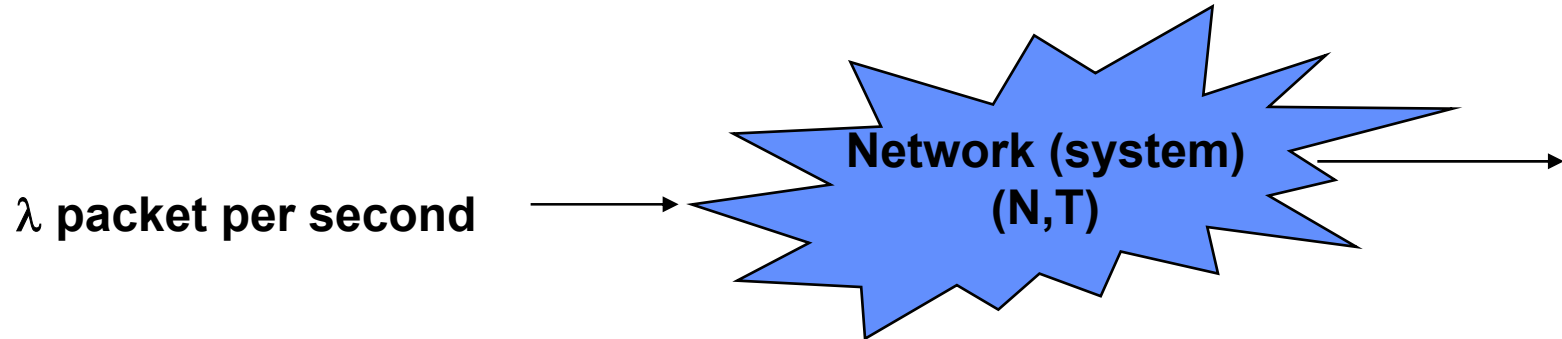
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- **Model for**
  - Customers waiting in line
  - Assembly line
  - Packets in a network (transmission line)
- **Want to know**
  - Average number of customers in the system
  - Average delay experienced by a customer
- **Quantities obtained in terms of**
  - Arrival rate of customers (average number of customers per unit time)
  - Service rate (average number of customers that the server can serve per unit time)

# Little's theorem

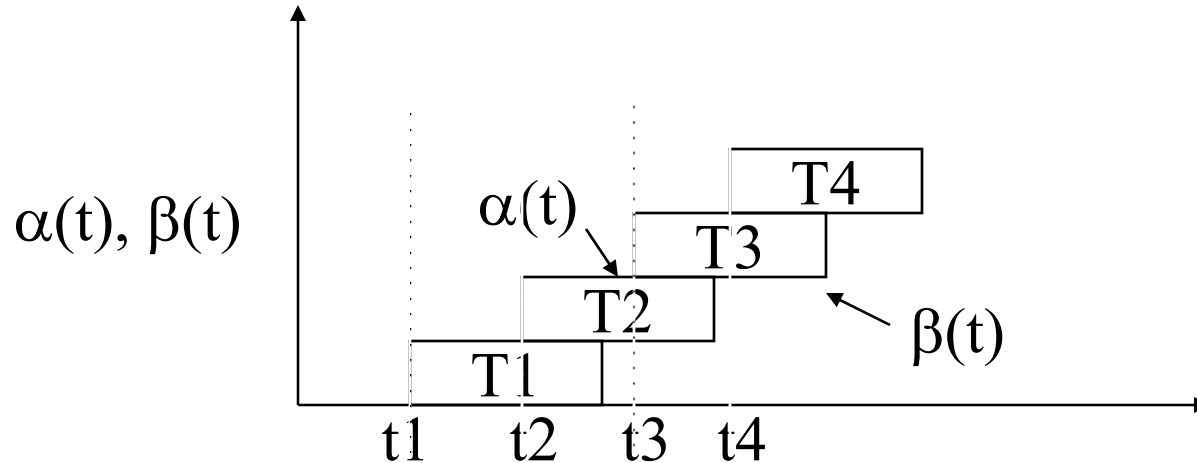
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- **N = average number of packets in system**
- **T = average amount of time a packet spends in the system**
- **$\lambda$  = arrival rate of packets into the system (not necessarily Poisson)**
  
- **Little's theorem:  $N = \lambda T$** 
  - Can be applied to entire system or any part of it
  - Crowded system -> long delays  
On a rainy day people drive slowly and roads are more congested!

# Proof of Little's Theorem

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- $\alpha(t)$  = number of arrivals by time  $t$
- $\beta(t)$  = number of departures by time  $t$
- $t_i$  = arrival time of  $i^{\text{th}}$  customer
- $T_i$  = amount of time  $i^{\text{th}}$  customer spends in the system
- $N(t)$  = number of customers in system at time  $t = \alpha(t) - \beta(t)$
  
- Similar proof for non First-come-first-serve

# Proof of Little's Theorem

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$$N_t = \frac{1}{t} \int_0^t N(\tau) d\tau = \text{timeave. number of customers in queue}$$

$$N = \text{Limit}_{t \rightarrow \infty} N_t = \text{steadystate time ave.}$$

$$\lambda_t = \alpha(t) / t, \quad \lambda = \text{Limit}_{t \rightarrow \infty} \lambda_t = \text{arrival rate}$$

$$T_t = \frac{\sum_{i=0}^{\alpha(t)} T_i}{\alpha(t)} = \text{timeave. system delay}, \quad T = \text{Limit}_{t \rightarrow \infty} T_t$$

- **Assume above limits exists, assume Ergodic system**

$$N(t) = \alpha(t) - \beta(t) \Rightarrow N_t = \frac{\sum_{i=1}^{\alpha(t)} T_i}{t}$$

$$N = \lim_{t \rightarrow \infty} \frac{\sum_{i=1}^{\alpha(t)} T_i}{t}, \quad T = \lim_{t \rightarrow \infty} \frac{\sum_{i=1}^{\alpha(t)} T_i}{\alpha(t)} \Rightarrow \sum_{i=1}^{\alpha(t)} T_i = \alpha(t) T$$

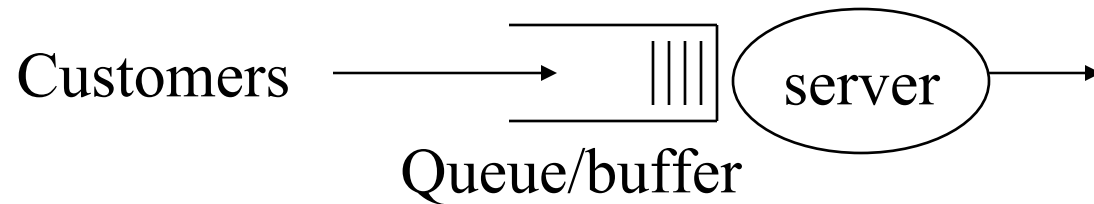
$$N = \frac{\sum_{i=1}^{\alpha(t)} T_i}{t} = \left( \frac{\alpha(t)}{t} \right) \frac{\sum_{i=1}^{\alpha(t)} T_i}{\alpha(t)} = \lambda T$$

# Application of little's Theorem

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- Little's Theorem can be applied to almost any system or part of it

- Example:



## 1) The transmitter: $D_{TP}$ = packet transmission time

- Average number of packets at transmitter =  $\lambda D_{TP} = \rho$  = link utilization

## 2) The transmission line: $D_p$ = propagation delay

- Average number of packets in flight =  $\lambda D_p$

## 3) The buffer: $D_q$ = average queueing delay

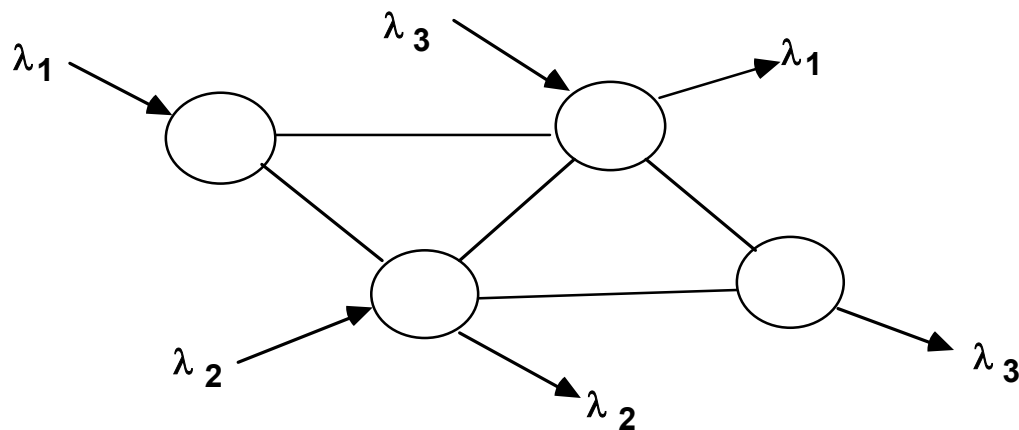
- Average number of packets in buffer =  $N_q = \lambda D_q$

## 4) Transmitter + buffer

- Average number of packets =  $\rho + N_q$

# Application to complex system

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- We have complex network with several traffic streams moving through it and interacting arbitrarily
- For each stream  $i$  individually, Little says  $N_i = \lambda_i T_i$
- For the streams collectively, Little says  $N = \lambda T$  where

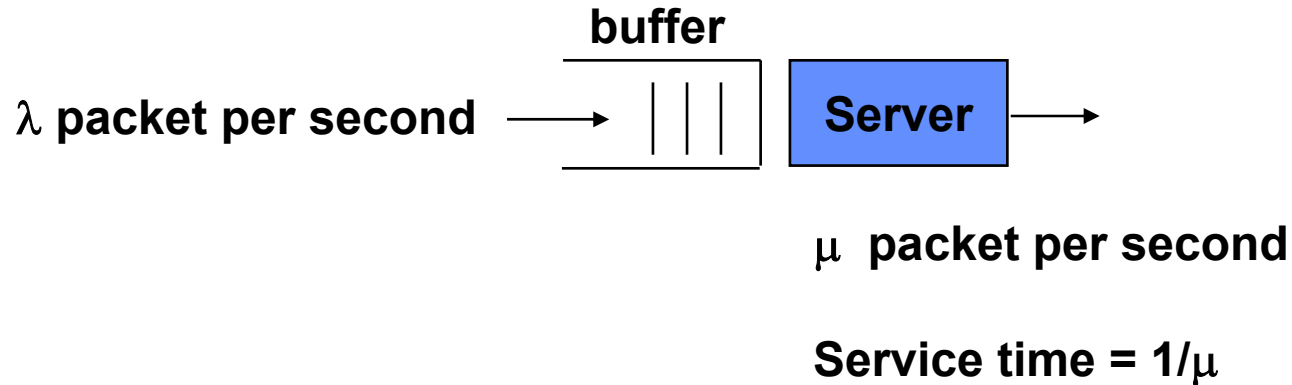
$$N = \sum_i N_i \quad \& \quad \lambda = \sum_i \lambda_i$$

From Little's Theorem:

$$T = \frac{\sum_{i=1}^{i=k} \lambda_i T_i}{\sum_{i=1}^{i=k} \lambda_i}$$

# Single server queues

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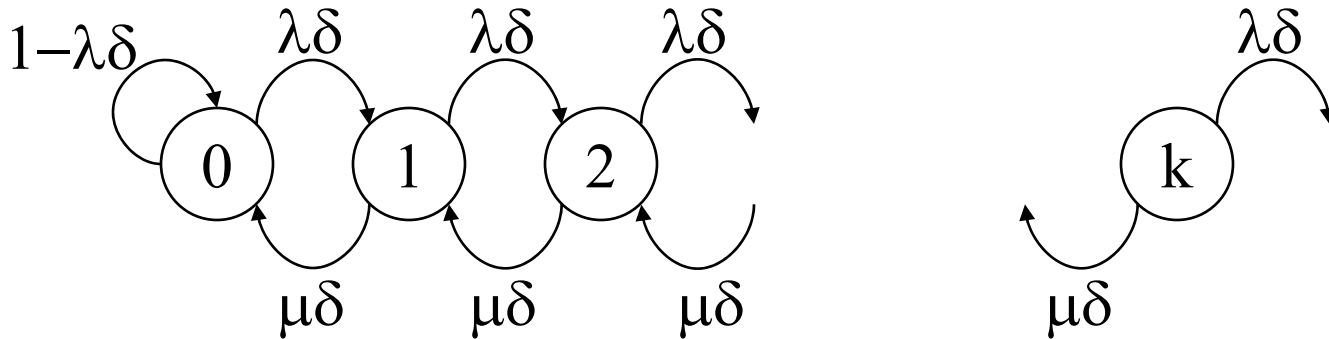


- **M/M/1**
  - Poisson arrivals, exponential service times
- **M/G/1**
  - Poisson arrivals, general service times
- **M/D/1**
  - Poisson arrivals, deterministic service times (fixed)



# Markov Chain for M/M/1 system

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- State  $k \Rightarrow k$  customers in the system
- $P(i,j)$  = probability of transition from state  $i$  to state  $j$ 
  - As  $\delta \Rightarrow 0$ , we get:

$$P(0,0) = 1 - \lambda\delta,$$

$$P(j,j) = 1 - \lambda\delta - \mu\delta$$

$$P(j,j+1) = \lambda\delta$$

$$P(j,j-1) = \mu\delta$$

$$P(i,j) = 0 \text{ for all other values of } i,j.$$

- Birth-death chain: Transitions exist only between adjacent states
  - $\lambda\delta, \mu\delta$  are flow rates between states

# Equilibrium analysis

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- We want to obtain  $P(n)$  = the probability of being in state  $n$
- At equilibrium  $\lambda P(n) = \mu P(n+1)$  for all  $n$ 
  - Local balance equations between two states  $(n, n+1)$
  - $P(n+1) = (\lambda/\mu)P(n) = \rho P(n)$ ,  $\rho = \lambda/\mu$

- It follows:  $P(n) = \rho^n P(0)$

$$\sum_{i=0}^{\infty} P(n) = 1$$

- Now by axiom of probability:

$$\Rightarrow \sum_{i=0}^{\infty} \rho^n P(0) = \frac{P(0)}{1 - \rho} = 1$$

$$\Rightarrow P(0) = 1 - \rho$$

$$P(n) = \rho^n (1 - \rho)$$

# Average queue size

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$$N = \sum_{n=0}^{\infty} nP(n) = \sum_{n=0}^{\infty} n\rho^n(1-\rho) = \frac{\rho}{1-\rho}$$

$$N = \frac{\rho}{1-\rho} = \frac{\lambda/\mu}{1-\lambda/\mu} = \frac{\lambda}{\mu-\lambda}$$

- **N = Average number of customers in the system**
- **The average amount of time that a customer spends in the system can be obtained from Little's formula ( $N=\lambda T \Rightarrow T = N/\lambda$ )**
- **T includes the queueing delay plus the service time (Service time =  $D_{TP} = 1/\mu$ )**
  - **W = amount of time spent in queue =  $T - 1/\mu \Rightarrow$**
- **Finally, the average number of customers in the buffer can be obtained from little's formula**

$$T = \frac{1}{\mu - \lambda}$$

$$W = \frac{1}{\mu - \lambda} - \frac{1}{\mu}$$

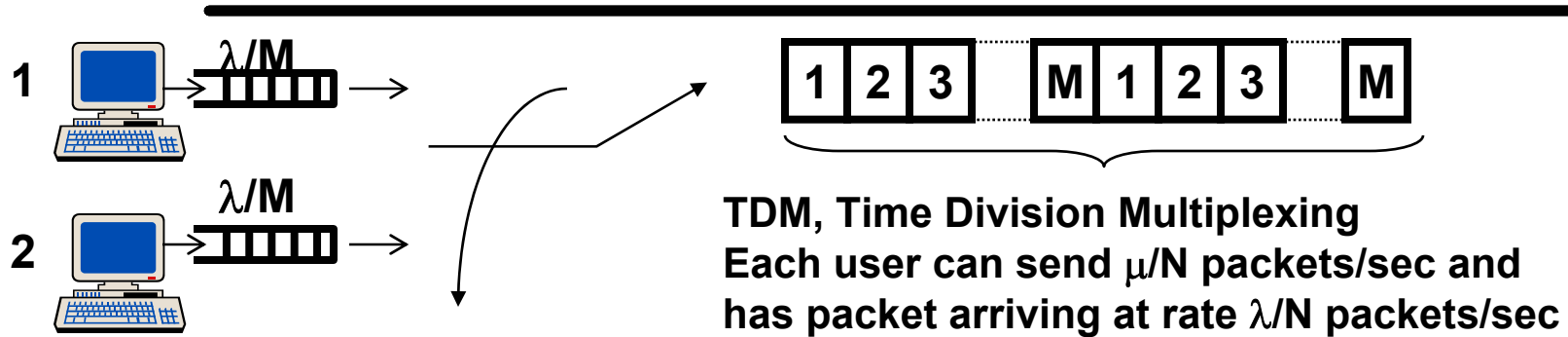
$$N_Q = \lambda W = \frac{\lambda}{\mu - \lambda} - \frac{\lambda}{\mu} = N - \rho$$

# Example (fast food restaurant)

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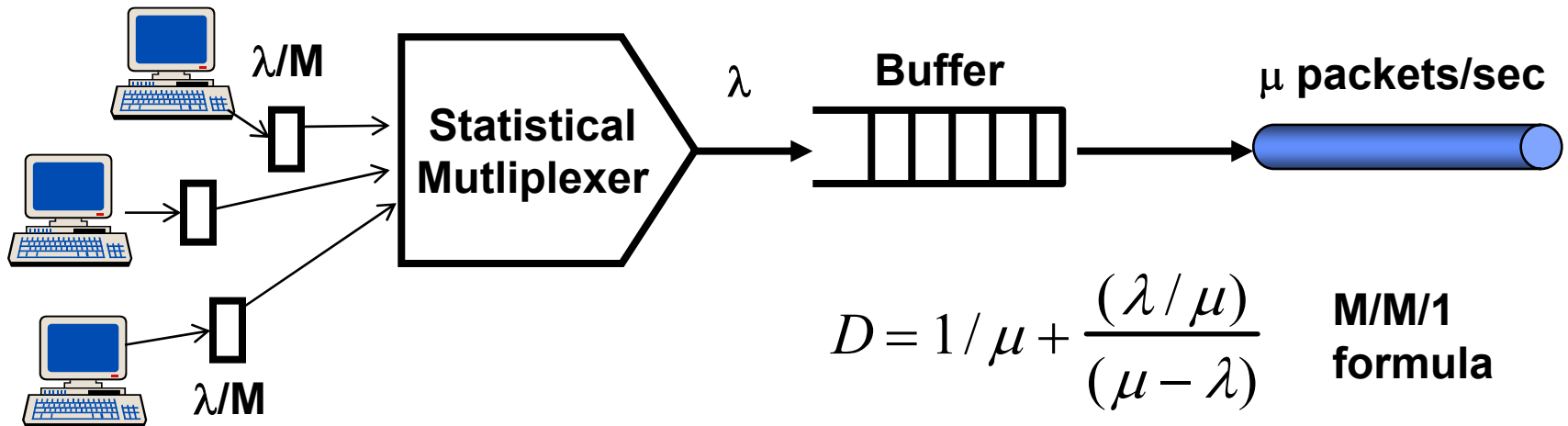
- Customers arrive at a fast food restaurant at a rate of 100 per hour and take 30 seconds to be served.
- How much time do they spend in the restaurant?
  - Service rate =  $\mu = 60/0.5 = 120$  customers per hour
  - $T = 1/(\mu - \lambda) = 1/(120 - 100) = 1/20$  hrs = 3 minutes
- How much time waiting in line?
  - $W = T - 1/\mu = 2.5$  minutes
- How many customers in the restaurant?
  - $N = \lambda T = 5$
- What is the server utilization?
  - $\rho = \lambda/\mu = 5/6$

# Packet switching vs. Circuit switching

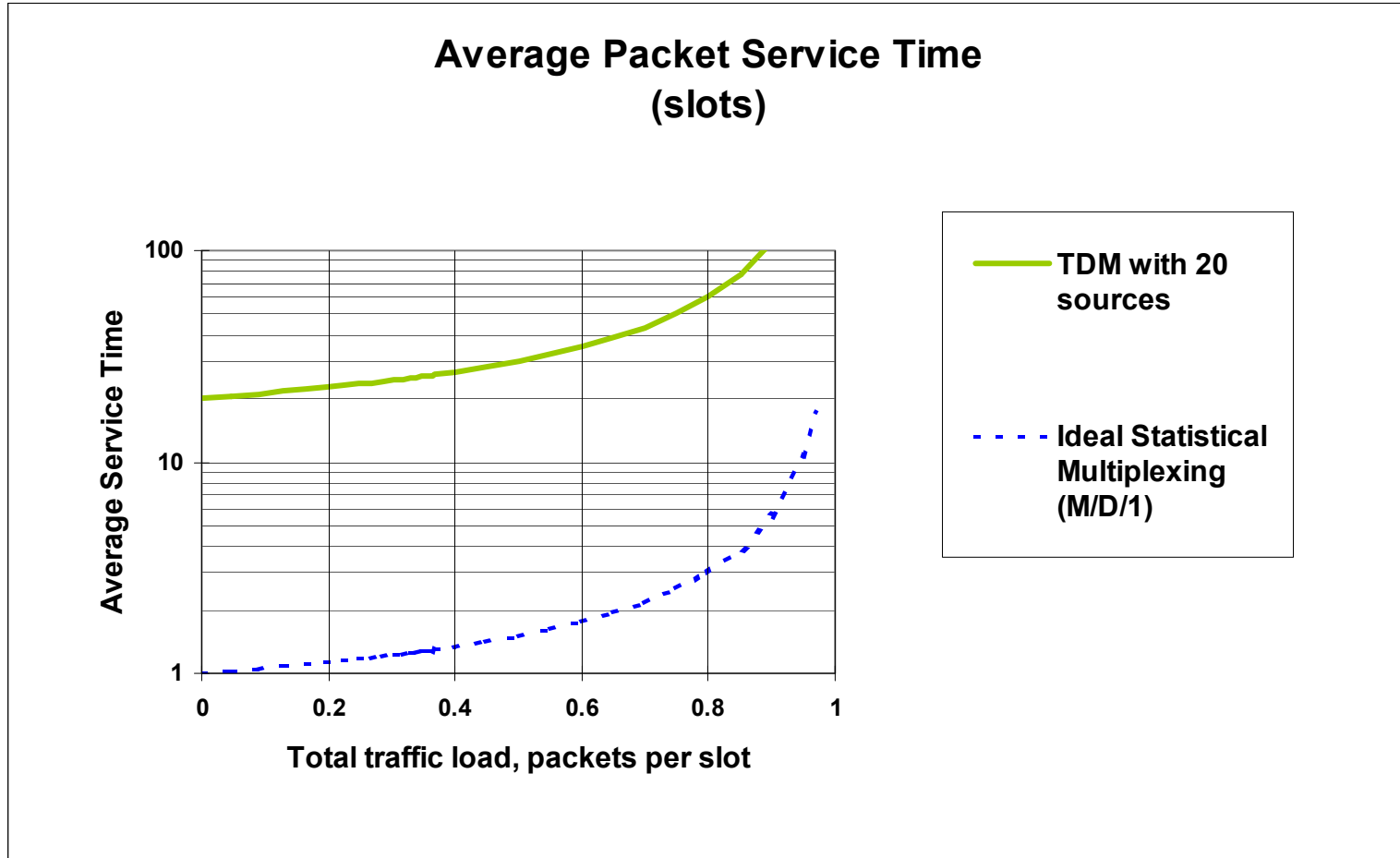


$$D = M / \mu + \frac{M(\lambda / \mu)}{(\mu - \lambda)} \quad \text{M/M/1 formula}$$

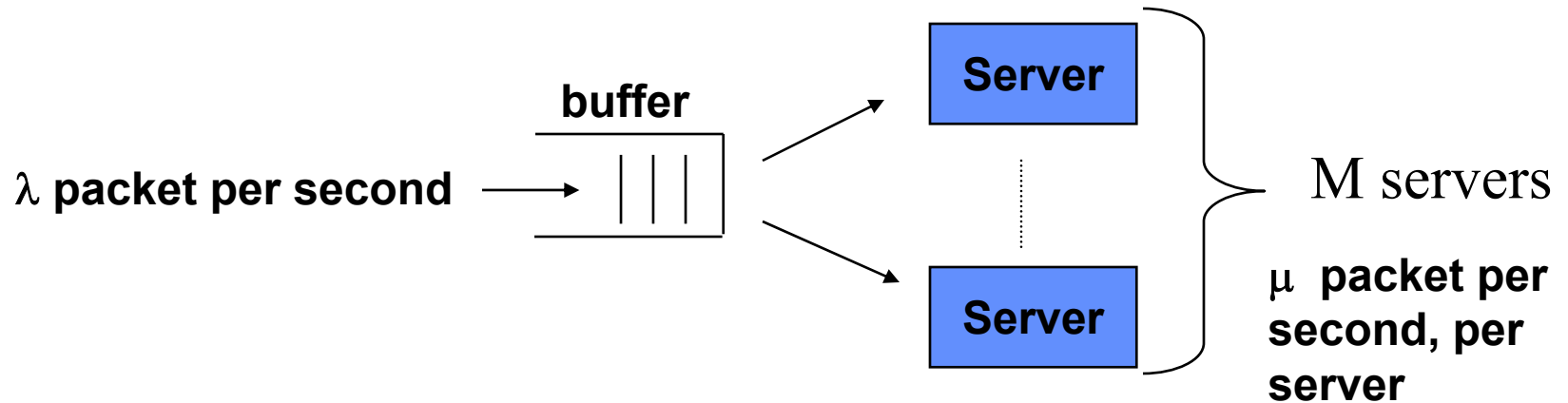
Packets generated at random times



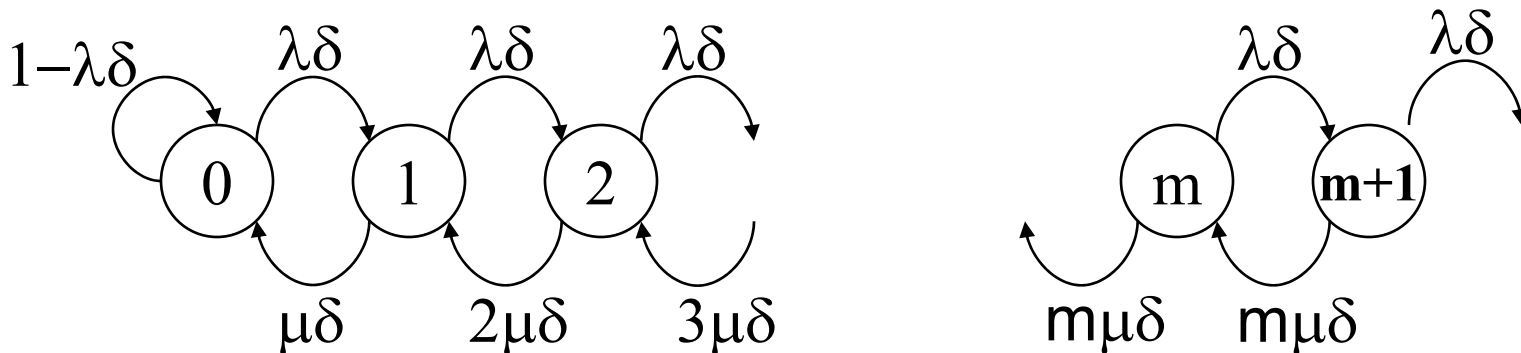
# Circuit (tdm/fdm) vs. Packet switching



# M server systems: M/M/m



- Departure rate is proportional to the number of servers in use
- Similar Markov chain:



# M/M/m queue

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- **Balance equations:**

$$\lambda P(n-1) = n\mu P(n) \quad n \leq m$$

$$\lambda P(n-1) = m\mu P(n) \quad n > m$$

$$P(n) = \begin{cases} P(0)(m\rho)^n / n! & n \leq m \\ P(0)(m^m \rho^n) / m! & n > m \end{cases} \quad \rho = \frac{\lambda}{m\mu} \leq 1$$

- **Again, solve for P(0):**

$$P(0) = \left[ \sum_{n=0}^{m-1} \frac{(m\rho)^n}{n!} + \frac{(m\rho)^m}{m!(1-\rho)} \right]^{-1}$$

$$P_Q = \sum_{n=m}^{\infty} P(n) = \frac{P(0)(m\rho)^m}{m!(1-\rho)}$$

$$N_Q = \sum_{n=0}^{\infty} nP(n+m) = \sum_{n=0}^{\infty} nP(0) \left( \frac{m^m \rho^{m+n}}{m!} \right) = P_Q \left( \frac{\rho}{1-\rho} \right)$$

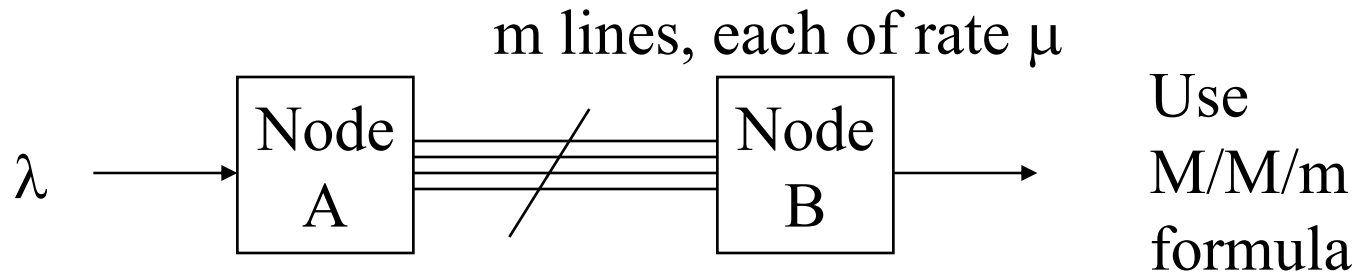
$$W = \frac{N_Q}{\lambda}, \quad T = W + 1/\mu, \quad N = \lambda T = \lambda/\mu + N_Q$$



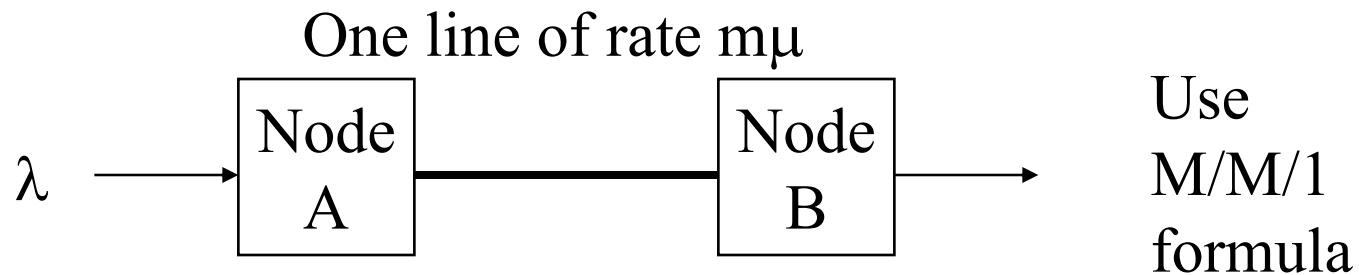
# Applications of M/M/m

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- Bank with  $m$  tellers
- Network with parallel transmission lines



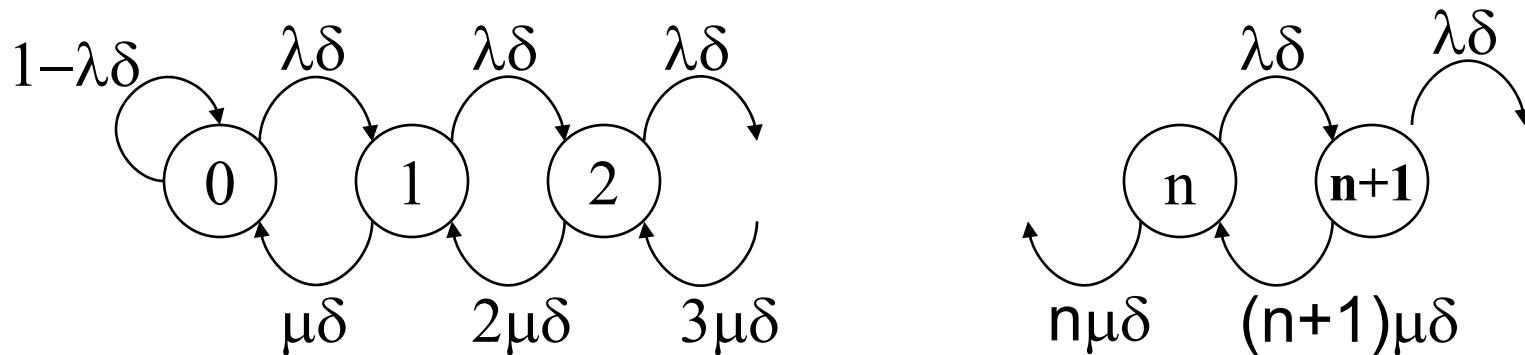
VS



- When the system is lightly loaded,  $PQ \sim 0$ , and Single server is  $m$  times faster
- When system is heavily loaded, queueing delay dominates and systems are roughly the same

# M/M/Infinity

- **Unlimited servers => customers experience no queueing delay**
- **The number of customers in the system represents the number of customers presently being served**



$$\lambda P(n-1) = n\mu P(n), \forall n > 1, \Rightarrow P(n) = \frac{P(0)(\lambda/\mu)^n}{n!}$$

$$P(0) = \left[ 1 + \sum_{n=1}^{\infty} (\lambda/\mu)^n / n! \right]^{-1} = e^{-\lambda/\mu}$$

$$P(n) = (\lambda/\mu)^n e^{-\lambda/\mu} / n! \Rightarrow \text{Poisson distribution!}$$

$$N = \text{Average number in system} = \lambda / \mu, T = N / \lambda = 1 / \mu = \text{service time}$$

# Blocking Probability

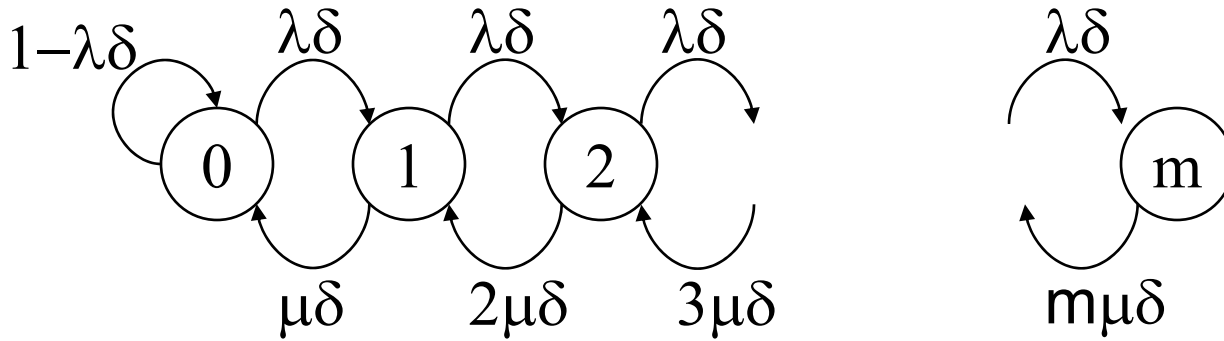
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- **A circuit switched network can be viewed as a Multi-server queueing system**
  - Calls are blocked when no servers available - “busy signal”
  - For circuit switched network we are interested in the call blocking probability
- **M/M/m/m system**
  - m servers => m circuits
  - Last m indicated that the system can hold no more than m users
- **Erlang B formula**
  - Gives the probability that a caller finds all circuits busy
  - Holds for general call arrival distribution (although we prove Markov case only)

$$P_B = \frac{(\lambda / \mu)^m / m!}{\sum_{n=0}^m (\lambda / \mu)^n / n!}$$

# M/M/m/m system: Erlang B formula

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$$\lambda P(n-1) = n\mu P(n), 1 \leq n \leq m, \Rightarrow P(n) = \frac{P(0)(\lambda/\mu)^n}{n!}$$

$$P(0) = \left[ \sum_{n=0}^m (\lambda/\mu)^n / n! \right]^{-1}$$

$$P_B = P(\text{Blocking}) = P(m) = \frac{(\lambda/\mu)^m / m!}{\sum_{n=0}^m (\lambda/\mu)^n / n!}$$

# Erlang B formula

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- **System load usually expressed in Erlangs**

- $A = \lambda/\mu = (\text{arrival rate}) * (\text{ave call duration}) = \text{average load}$
- Formula insensitive to  $\lambda$  and  $\mu$  but only to their ratio

$$P_B = \frac{(A)^m / m!}{\sum_{n=0}^m (A)^n / n!}$$

- **Used for sizing transmission line**

- How many circuits does the satellite need to support?
- The number of circuits is a function of the blocking probability that we can tolerate  
Systems are designed for a given load predictions and blocking probabilities (typically small)

- **Example**

- Arrival rate = 4 calls per minute, average 3 minutes per call =>  $A = 12$
- How many circuits do we need to provision?  
Depends on the blocking probability that we can tolerate

<u>Circuits</u>	<u><math>P_B</math></u>
20	1%
15	8%
7	30%

# Multi-dimensional Markov Chains

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- **K classes of customers**
  - Class  $j$ : arrival rate  $\lambda_j$ ; service rate  $\mu_j$
- **State of system:  $n = (n_1, n_2, \dots, n_k)$ ;  $n_j$  = number of class  $j$  customers in the system**
- **If detailed balance equations hold for adjacent states, then a product form solution exists, where:**
  - $P(n_1, n_2, \dots, n_k) = P_1(n_1) * P_2(n_2) * \dots * P_k(n_k)$
- **Example: K independent M/M/1 systems**

$$P_i(n_i) = \rho_i^{n_i} (1 - \rho_i), \quad \rho_i = \lambda_i / \mu_i$$

- **Same holds for other independent birth-death chains**
  - E.g., M.M/m, M/M/Inf, M/M/m/m

# Truncation

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- **Eliminate some of the states**
  - E.g., for the K M/M/1 queues, eliminate all states where  $n_1+n_2+\dots+n_k > K_1$  (some constant)
- **Resulting chain must remain irreducible**
  - All states must communicate

## Product form for stationary distribution of the truncated system

- E.g., K independent M/M/1 queues

$$P(n_1, n_2, \dots, n_k) = \frac{\rho_1^{n_1} \rho_2^{n_2} \dots \rho_K^{n_K}}{G}, \quad G = \sum_{n \in S} \rho_1^{n_1} \rho_2^{n_2} \dots \rho_K^{n_K}$$

- E.g., K independent M/M/inf queues

$$P(n_1, n_2, \dots, n_k) = \frac{(\rho_1^{n_1} / n_1!) (\rho_2^{n_2} / n_2!) \dots (\rho_K^{n_K} / n_k!)}{G}, \quad G = \sum_{n \in S} (\rho_1^{n_1} / n_1!) (\rho_2^{n_2} / n_2!) \dots (\rho_K^{n_K} / n_k!)$$

- **G** is a normalization constant that makes **P(n)** a distribution
- **S** is the set of states in the truncated system

# Example

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- **Two session classes in a circuit switched system**
  - M channels of equal capacity
  - Two session types:
    - Type 1: arrival rate  $\lambda_1$  and service rate  $\mu_1$
    - Type 2: arrival rate  $\lambda_2$  and service rate  $\mu_2$
- **System can support up to M sessions of either class**
  - If  $\mu_1 = \mu_2$ , treat system as an M/M/m/m queue with arrival rate  $\lambda_1 + \lambda_2$
  - When  $\mu_1 \neq \mu_2$  need to know the number of calls in progress of each session type
  - Two dimensional markov chain state =  $(n_1, n_2)$
  - Want  $P(n_1, n_2)$ :  $n_1 + n_2 \leq m$
- **Can be viewed as truncated M/M/Inf queues**
  - Notice that the transition rates in the M/M/Inf queue are the same as those in a truncated M/M/m/m queue

$$P(n_1, n_2) = \frac{(\rho_1^{n_1} / n_1!)(\rho_2^{n_2} / n_2!)}{G}, \quad G = \sum_{i=0}^m \sum_{j=0}^{m-i} (\rho_1^i / i!)(\rho_2^j / j!), \quad n_1 + n_2 \leq m$$

- Notice that the double sum counts only states for which  $j+i \leq m$



# PASTA: Poisson Arrivals See Time Averages

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- The state of an M/M/1 queue is the number of customers in the system
- More general queueing systems have a more general state that may include how much service each customer has already received
- For Poisson arrivals, the arrivals in any future increment of time is independent of those in past increments and for many systems of interest, independent of the present state  $S(t)$  (true for M/M/1, M/M/m, and M/G/1).
- For such systems,  $P\{S(t)=s|A(t+\delta)-A(t)=1\} = P\{S(t)=s\}$ 
  - (where  $A(t)$  = # arrivals since  $t=0$ )
- In steady state, arrivals see steady state probabilities

# Occupancy distribution upon arrival

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- Arrivals may not always see the steady-state averages
- Example:
  - Deterministic arrivals 1 per second
  - Deterministic service time of 3/4 seconds

$\lambda = 1$  packets/second  $T = 3/4$  seconds (no queueing)

$$N = \lambda T = \text{Average occupancy} = 3/4$$

- However, notice that an arrival always finds the system empty!

# Occupancy upon arrival for a M/M/1 queue

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$$a_n = \lim_{t \rightarrow \infty} (P(N(t) = n \mid \text{an arrival occurred just after time } t))$$

$$P_n = \lim_{t \rightarrow \infty} (P(N(t) = n))$$

For M/M/1 systems  $a_n = P_n$

**Proof:** Let  $A(t, t+\delta)$  be the event that an arrival occurred between  $t$  and  $t+\delta$

$$\begin{aligned} a_n(t) &= \lim_{t \rightarrow \infty} (P(N(t) = n \mid A(t, t+\delta))) \\ &= \lim_{t \rightarrow \infty} (P(N(t) = n, A(t, t+\delta)) / P(A(t, t+\delta))) \\ &= \lim_{t \rightarrow \infty} P(A(t, t+\delta) \mid N(t) = n) P(N(t) = n) / P(A(t, t+\delta)) \end{aligned}$$

- Since future arrivals are independent of the state of the system,

$$P(A(t, t+\delta) \mid N(t) = n) = P(A(t, t+\delta))$$

- Hence,  $a_n(t) = P(N(t) = n) = P_n(t)$
- Taking limits as  $t \rightarrow \infty$ , we obtain  $a_n = P_n$
- Result holds for M/G/1 systems as well